

Combining diverse information sources with the II-CC-FF paradigm, with applications in meta-analysis and beyond



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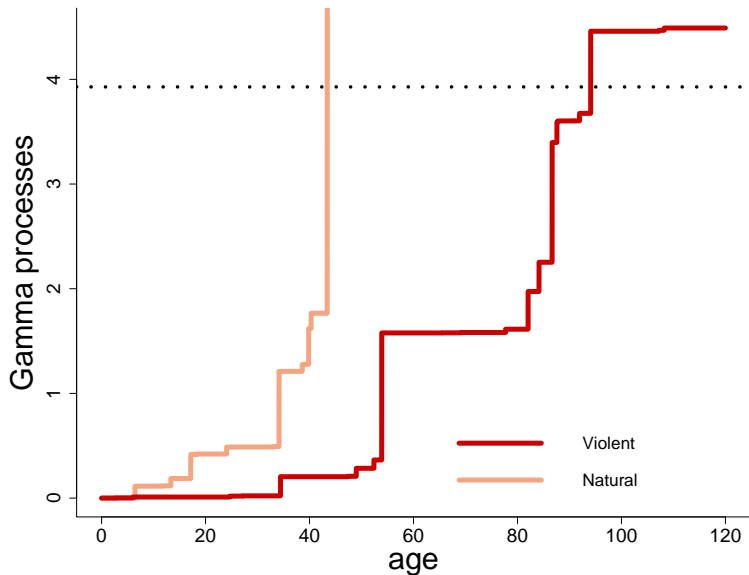
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First, something about my other projects

- Cunen, C., Hermansen, G. and Hjort, N.L. (2017). **Confidence distributions for change-points and regime shifts.** *Journal of Statistical Planning and Inference.*
- Model selection for meta-analysis of 2×2 tables using the focused information criteria (FIC) – in progress (with NLH).
- A FIC for linear mixed effect (LME) models:
 - An application: Cunen, C., Walløe, L. and Hjort, N.L. (2017). **Decline in energy storage in Antarctic Minke whales during the JARPA period: Assessment via the Focused Information Criterion (FIC).** Reports of the Scientific Committee of the International Whaling Commission.
 - A "methods" paper on its way: **Model selection for linear mixed models via the Focused Information Criterion, with an application to whale ecology.**
- Competing risks with gamma process threshold crossing models (with NLH).

Gamma process models for competing risks



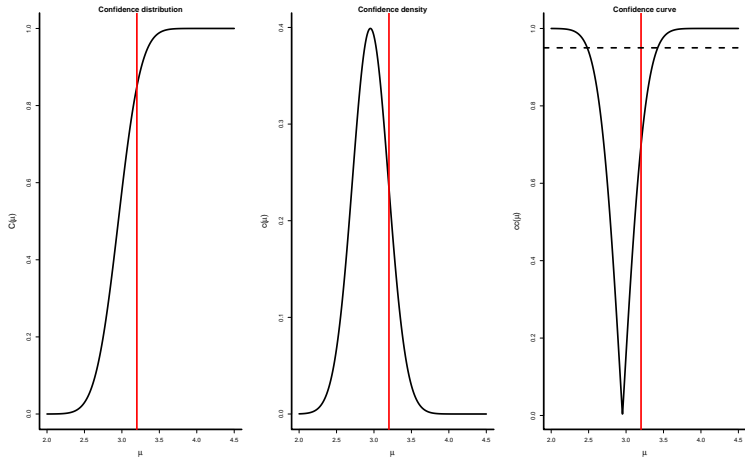
The Problem - Combination of information

We have **independent** data sources $1, \dots, k$ providing information about parameters ψ_1, \dots, ψ_k .

Our interest is in the overall focus parameter $\phi = \phi(\psi_1, \dots, \psi_k)$.

- II-CC-FF: a general framework to provide inference for ϕ in cases like this.
 - Similar to *likelihood synthesis* from Schweder and Hjort (1996)
- Beyond ordinary meta-analysis:
 - not restricted to cases where the sources inform on the *same* parameter – we can deal with complex functions of the parameters from each source: $\phi = \phi(\psi_1, \dots, \psi_k)$
 - we can deal with cases where we only have summary statistics from some or all of the sources
 - we can handle very diverse sources – for example combining parametric and non-parametric analyses

Confidence distributions (CD)



- \approx a posterior without having to specify a prior
 - a sample-dependent distribution function on the parameter space
 - can be used for inference (for example for constructing confidence intervals of all levels)

Requirements for CDs

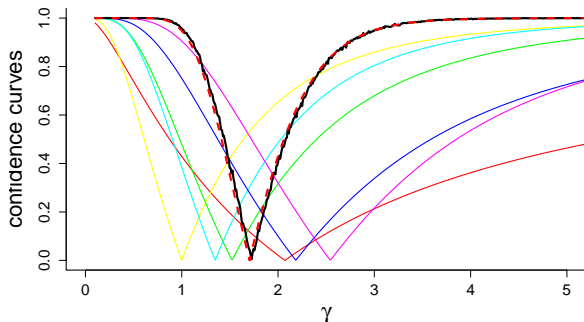
Definition

A function $C(\theta, Y)$ is called a confidence distribution for a parameter θ if:

- $C(\theta, Y)$ is a cumulative distribution function on the parameter space
- at the true parameter value $\theta = \theta_0$, $C(\theta_0, Y)$ as a function of the random sample Y follows the uniform distribution $U[0,1]$
- The second requirement ensures that all confidence intervals have the correct coverage.
- Note that *any* method producing confidence intervals fulfilling these requirements can be used to make CDs (no-matter the underlying paradigm).
- More on CDs in *Confidence, Likelihood, Probability*. (Schweder and Hjort, 2016.)

Outline

- II-CC-FF - general procedure
- Examples illustrating what II-CC-FF can do
 - Classic meta-analysis
 - More complex meta-analysis: Blood loss
 - Random effects: All Blacks
 - Very diverse sources: First word



II-CC-FF - overview

Combining information, for inference about a **focus parameter**

$\phi = \phi(\psi_1, \dots, \psi_k)$:

II: Independent Inspection: From data source y_i to estimates and intervals, in the form of a confidence distribution/curve:

$$y_i \implies C_i(\psi_i)$$

CC: Confidence Conversion: From the confidence distribution to a confidence log-likelihood,

$$C_i(\psi_i) \implies \ell_{c,i}(\psi_i)$$

FF: Focused Fusion: Use the combined confidence log-likelihood $\ell_f(\psi_1, \dots, \psi_k) = \sum_{i=1}^k \ell_{c,i}(\psi_i)$ to construct a CD for the given focus $\phi = \phi(\psi_1, \dots, \psi_k)$, often via profiling:

$$\ell_f(\psi_1, \dots, \psi_k) \implies C_{fusion}(\phi)$$

Constructing CDs via Wilks' theorem

In many regular cases we have, at the true parameter value ψ_0 :

$$2\{\ell_{n,prof}(\hat{\psi}) - \ell_{n,prof}(\psi_0)\} \rightarrow_d \chi_1^2,$$

as the sample size n increases. This gives us our favourite approximate confidence curve construction,

$$cc(\psi) = \Gamma_1(2\{\ell_{n,prof}(\hat{\psi}) - \ell_{n,prof}(\psi)\}).$$

CC - Confidence Conversion

The most difficult step?

$$C_i(\psi_i) \implies \ell_{c,i}(\psi_i)$$

In some cases we will already have a log-likelihood for ψ_i from the **II**-step and then there are no problems.

In other cases, the confidence curves from the **II**-step are not constructed via likelihoods.

Then we need to do something else (and be more careful). A simple and general method - **the normal conversion**:

$$\ell_c(\psi) = -\frac{1}{2}\Gamma_1^{-1}(\text{cc}(\psi, y)) = -\frac{1}{2}\{\Phi^{-1}(C(\psi, y))\}^2.$$

[Note that the confidence log-likelihood is not equal to the log-confidence density ($\log \partial C(\psi)/\partial \psi$). Except under the normal model.]

We can deal with 1: Classic meta-analysis

Assume all sources inform on the exact same parameter

$\psi_1 = \dots = \psi_k = \psi$, and that each source provide estimators $\hat{\psi}_i$ that are normally distributed $N(\psi, \sigma_i^2)$ with known σ_i s.

II: Data source y_i leads to $C_i(\psi) = \Phi((\psi - \hat{\psi}_i)/\sigma_i)$.

CC: From $C_i(\psi)$ to $\ell_{c,i}(\psi) = -\frac{1}{2}(\psi - \hat{\psi}_i)^2/\sigma_i^2$.

FF: Summing $\ell_f(\psi) = \sum_{i=1}^k \ell_{c,i}(\psi)$ leads to the classic answer

$$\hat{\psi} = \frac{\sum_{i=1}^k \hat{\psi}_i / \sigma_i^2}{\sum_{i=1}^k 1 / \sigma_i^2} \sim N \left(\psi, \left(\sum_{i=1}^k 1 / \sigma_i^2 \right)^{-1} \right).$$

We can deal with 2: more complicated meta-analysis

We do not have access to the full dataset (only summaries), and the studies differ in their reported outcomes: some studies report continuous outcomes, others report counts of a binary outcome.

Example from Whitehead et al. (1999): **Blood loss during labor**. Does treatment with oxytocic drugs help reduce blood loss?

Total of 11 studies, 6 studies report summary statistics of the continuous outcome (the actual blood loss in ml):

	Treatment	<i>n</i>	Mean	SD
Study 1	Control	510	325.75	288.61
	Treatment	490	255.60	213.74
Study 2

5 studies report counts of a binary outcome (yes = blood loss greater than 500 ml):

	Treatment	Yes	No	Total
Study 7	Control	152	697	849
	Treatment	50	796	846
Study 8

2: Blood loss

Model: $y_{ij} = \alpha_i + \beta z_{ij} + \epsilon_{ij}$ $\epsilon_{ij} \sim N(0, \sigma^2)$

II and **CC**: If study i has a continuous outcome we have

$$\ell_{c,i}(\alpha_i, \beta, \sigma) = -(n_1 + n_2) \log(\sigma) - \{(n_1 - 1)s_1^2 + (n_2 - 1)s_2^2 + n_1(\bar{y}_1 - \alpha_i)^2 + n_2(\bar{y}_2 - \alpha_i - \beta)^2\} / (2\sigma^2).$$

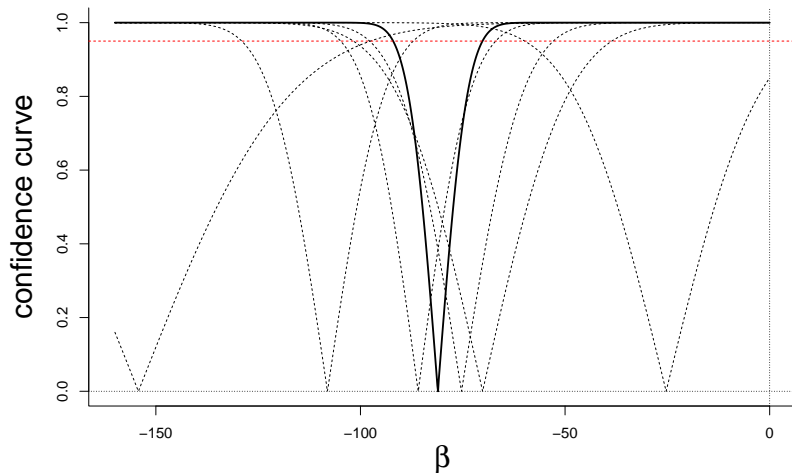
If study i has a binary outcome we have $\ell_{b,i}(\alpha_i, \beta, \sigma) = x_{01} \log\{\Phi((500 - \alpha_i)/\sigma)\} + x_{11} \log\{1 - \Phi((500 - \alpha_i)/\sigma)\} + x_{02} \log\{\Phi((500 - \alpha_i - \beta)/\sigma)\} + x_{12} \log\{1 - \Phi((500 - \alpha_i - \beta)/\sigma)\}.$

FF: Summing

$\ell_f(\alpha_1, \dots, \alpha_k, \beta, \sigma) = \sum_{i=1}^{k_c} \ell_{c,i}(\alpha_i, \beta, \sigma) + \sum_{i=1}^{k_b} \ell_{b,i}(\alpha_i, \beta, \sigma)$ and then profiling $\ell_{f,prof}(\beta) = \ell_f(\hat{\alpha}_1(\beta), \dots, \hat{\alpha}_k(\beta), \beta, \hat{\sigma}(\beta))$ and we get a combined confidence curve for β by using the Wilks' approximation

$$cc(\beta) = \Gamma_1\{2(\ell_{f,prof}(\hat{\beta}) - \ell_{f,prof}(\beta))\}.$$

2: Blood loss



Oxytocic drugs is seen to decrease blood loss.

We can deal with 3: Random effects!



3: All Blacks

We have measures of “passage times” in 10 Rugby games (\approx studies). There are 5 games before a certain change of rules and 5 after.

Model: $y_{ij} \sim \text{Gamma}(a_i, b_i)$

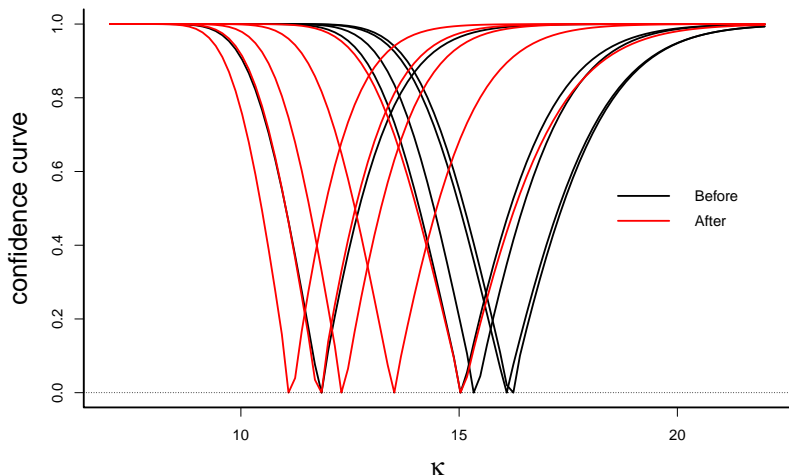
Say we are interested in the standard deviation of passage times

$$\kappa_i = \sqrt{a_i/b_i}.$$

It is relatively straightforward to construct confidence curves for each κ_i (by profiling and Wilks' approximation).

$$\text{II: } \text{cc}_i(\kappa_i) = \Gamma_1\{2(\ell_{i,\text{prof}}(\hat{\kappa}_i) - \ell_{i,\text{prof}}(\kappa_i))\}$$

3: All Blacks



There seems to be smaller standard deviations in passage times (smaller κ s) after the rule change.

But there is substantial spread in the mean κ s between different games - do we need random effects?

3: All Blacks

We assume :

- Before rule change $\kappa_1, \dots, \kappa_5 \sim N(\kappa_B, \tau_B^2)$ and

- After rule change $\kappa_6, \dots, \kappa_{10} \sim N(\kappa_A, \tau_A^2)$.

We are interested in making confidence curves for κ_B , κ_A , and the ratio between them $\delta = \kappa_B / \kappa_A$.

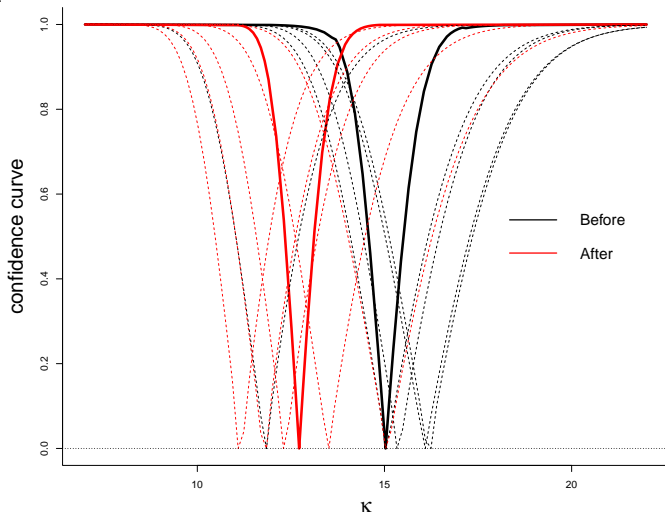
CC: We already have the $\ell_{i,prof}(\kappa_i)$ from the II-step, but now we need $\ell_i(\kappa_B, \tau_B)$ (and similarly for the parameters after the rule change):

$$\ell_i(\kappa_B, \tau_B) = \log \left[\int \exp \{ \ell_{i,prof}(\kappa_i) - \ell_{i,prof}(\hat{\kappa}_i) \} \frac{1}{\tau_B} \phi \left(\frac{\kappa_i - \kappa_B}{\tau_B} \right) d\kappa_i \right]$$

(We integrate numerically or use Laplace approximation)

3: All Blacks

FF: Summing $\ell_f(\kappa_B, \tau_B) = \sum_{i=1}^{k_B} \ell_i(\kappa_B, \tau_B)$, profile to get $\ell_{f,prof}(\kappa_B)$ and Wilks' approximation to get $cc(\kappa_B)$. Similarly for $cc(\kappa_A)$.



3: All Blacks - computation

For fast computation of the Laplace approximation and the subsequent profiling, we can make use of the TMB package (Template Model Builder, related to the ADMB).

<https://github.com/kaskr/adcomp/wiki>

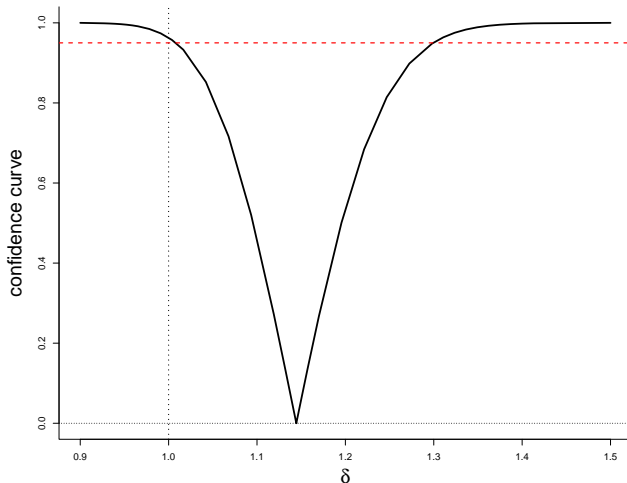
One simply(?) needs to write down an objective function/likelihood in C++. Then one specifies which parameters to profile out and which to integrate over.

3: All Blacks

The ratio $\delta = \kappa_B / \kappa_A$. Sum, profile and Wilks':

$\ell_{ff}(\kappa_B, \delta) = \ell_{f,prof,B}(\kappa_B) + \ell_{f,prof,A}(\kappa_B/\delta)$, then

$\ell_{ff}(\delta) = \ell_{ff}(\hat{\kappa}_B(\delta), \delta)$.



We can deal with 4: Very diverse sources

We have two sources:

- A **large study**: 1640 parents report the age (in months) at which their child said its first word. Ranges from 1 (!) to 25.
- A **small study**: 51 parents report the age (in months) at which their child said its first word. Here we have some covariate information: gender (of the child).

Focus: When do girls start to speak? And when do boys start to speak?

Model: proportional hazards model (no censoring here - but we could have dealt with that too!)

Question: can the large (low-quality) study improve our analysis of the small study?

Data from: Schneider, Yurovsky & Frank (2015). Large-scale investigations of variability in children's first words.

In CogSci2015 Proceedings.

4: Age at first word

Focus: probability that a child with covariate information x_0 does not speak at the age of 12 months

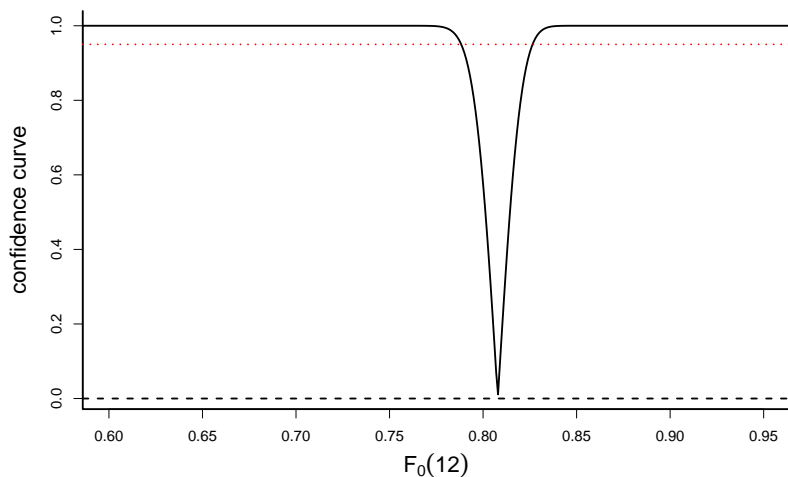
$$S(t_0|x_0) = e^{-H_0(t_0)e^{x_0^t\beta}} = S_0(t_0)e^{x_0^t\beta} = (1-F_0(t_0))e^{x_0^t\beta} \quad \text{with } t_0 = 12.$$

- **Large study:** will give information about F_0 at t_0 -
 $\implies \text{cc}_1(F_0(t_0))$. Non-parametric!
- **Small study:** will give information about $\beta \implies \text{cc}_2(\beta)$. Cox model - Semi-parametric!

with II-CC-FF we can combine these and obtain a cc for $S(t_0|x_0)$ with $t_0 = 12$.

4: Age at first word

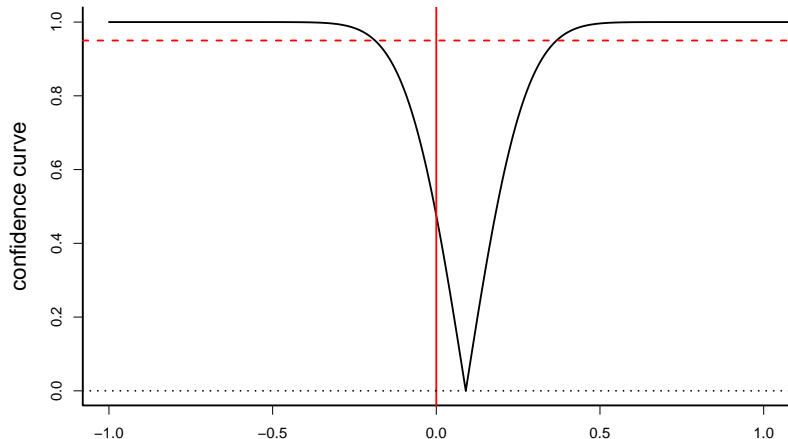
Obtaining a confidence curve for the “baseline” F_0 (at $t_0 = 12$).
An exact CD based on the binomial distribution.



4: Age at first word

Obtaining a confidence curve for the coefficient β (taking care to define gender as 1/-1, so that the value 0 corresponds to the overall mean)

Approximate CD based on the normal distribution (here we only need the summary statistics: estimate and standard error).

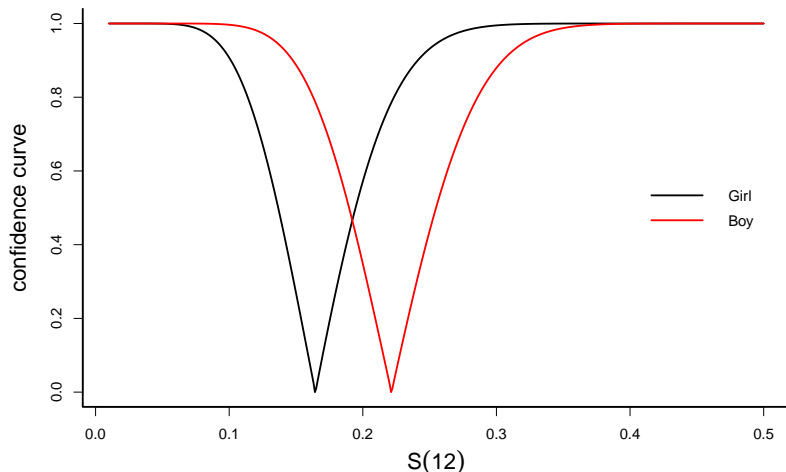


4: Age at first word

FF: summing $\ell_f(F_0, \beta) = \ell_1(F_0) + \ell_2(\beta)$ and profiling

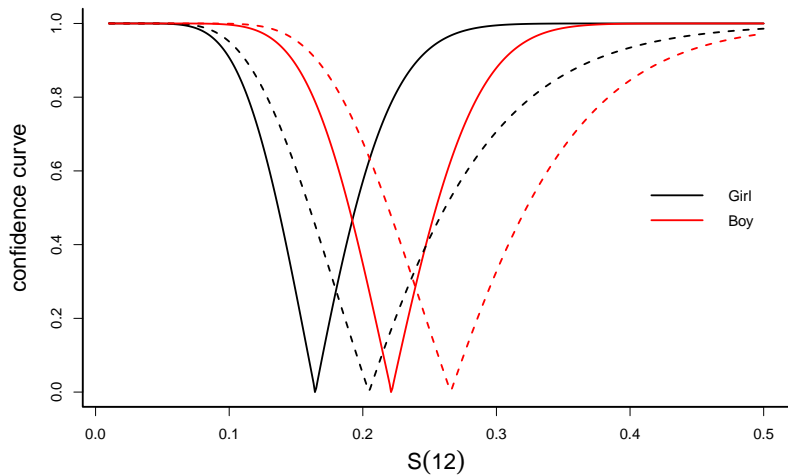
$$\ell_{f,prof}(S(t_0|x_0)) = \max\{\ell_f(F_0, \beta) : (1 - F_0(t_0))^{e^{x_0^t \beta}} = S(t_0|x_0)\}$$

and then Wilks' approximation.



4: Age at first word

Comparing with results from small source only.



Concluding remarks

What can II-CC-FF do?

- Deal with summary statistics
- Deal with complex functions of the parameters from each source: $\phi = \phi(\psi_1, \dots, \psi_k)$
- Deal with very diverse sources (hard and soft data; improving inference from small, carefully designed studies by using large, "less-informative" datasets; ...)

Some remaining challenges:

- Neyman-Scott type problems
- Corrections for improving the Wilks' approximation
- Real applications!

References

- Cunen and Hjort (2016). Combining information across diverse sources: The II-CC-FF paradigm. *Proceedings from the Joint Statistical Meeting 2016, the American Statistical Association*.
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