Correcting for covariate measurement error using regression calibration

Centre for Statistical Methodology Forum

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The effects of covariate measurement error

Regression calibration

Our model of interest

- Suppose we are interested in estimating the relationship between an outcome Y and covariates/explanatory variables X₁, ..., X_p.
- We might do this by fitting a regression model (e.g. linear regression, logistic regression, Cox proportional hazards model).
- e.g. linear regression:

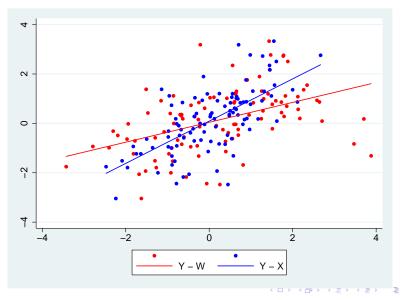
$$Y = \alpha + \beta_1 X_1 + \dots + \beta_p X_p + \epsilon$$

where $\epsilon \stackrel{\text{iid}}{\sim} N(0, \sigma_{\epsilon}^2)$.

Covariate measurement error

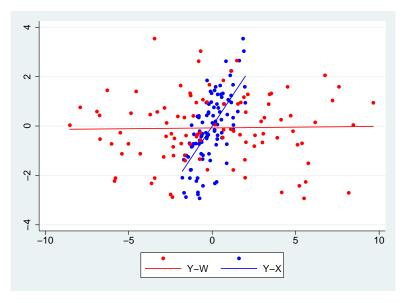
- For some explanatory variables, we may only have a noisy error-prone measurement W of the true exposure/confounder of interest.
- e.g. we have single measurements of blood pressure for each subject, but we are interested in X =subject's 'true' blood pressure at entry to our study.
- What effects does such noise/error have?
- One might think that completely random noise may reduce precision, but not bias our estimates.

The effects of covariate measurement error in linear regression



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More extreme measurement error



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The effects of covariate measurement error in linear regression

We will continue in this simplified setting:

$$Y = \alpha + \beta X + \epsilon$$

- Suppose that we actually measure W = X + U, where U is an independent measurement error.
- This is called the classical measurement error model.
- What slope β* do we estimate if we fit the model with W as the covariate, instead of X?
- One can show:

$$\beta^* = \beta \frac{Var(X)}{Var(X) + Var(U)}$$

The effects of covariate measurement error in linear regression

$$eta^* = eta rac{Var(X)}{Var(X) + Var(U)}$$

- The fraction on the right is what is referred to as the reliability of the measurements.
- As the amount of noise/error (Var(U)) increases, the slope is attenuated by a larger amount.
- The residual error in the regression also gets larger, so we have reduced power to detect associations.
- With multiple covariates, biases can be both towards no association (attenuation) or towards larger association.

Implications for epidemiology

- In the absence of confounders, exposure to disease associations are underestimated in magnitude.
- When we have confounders in our model, some of which are measured with error, our exposure to disease associations will not be properly adjusted for the confounders.
- When trying to determine which variables are most important in determining an outcome, variables can seem less important simply because they are measured imprecisely / with lots of error.

Allowing for covariate measurement error

- There are a number of statistical approaches to allowing for the effects of covariate measurement error.
- In addition to some assumptions about the nature of the errors (e.g. independence), all of them require some information on the magnitude of the errors.
- To quantify the magnitude of errors, we either need:
 - ► (Some) subjects to have repeated error-prone measurements, e.g. W₁ and W₂, or
 - ► (Some) subjects to have the true X measured, in addition to W.
- We will focus on the first situation.
- Such data allow us to estimate the magnitude of errors (Var(U)) and the variability of the true covariate (Var(X)).



The effects of covariate measurement error

Regression calibration

Regression calibration

- Regression calibration (RC) is one statistical approach for correcting for the effects of covariate measurement error.
- Rather than using W as the covariate in our model, we use E(X|W) as covariate.
- ► E(X|W) is the predicted value of X based on the error-prone measurement W.
- For classical error, $E(X|W) \neq W!$

Why RC works

As before, suppose:

$$Y = \alpha + \beta X + \epsilon$$

► Then taking expectations conditional on *W*:

$$E(Y|W) = \alpha + \beta E(X|W),$$

provided ϵ is independent of W (the non-differential error assumption).

This result means that if we regress Y on with E(X|W) as covariate, we obtain unbiased estimates of β.

Implementing RC consists of the following steps:

- 1. assume a model for X and W,
- 2. fit this model using data (usually from the study in question) where some subjects have repeat error-prone measurements,
- 3. calculate $\widehat{E}(X|W)$ based on the estimated parameters,
- 4. run the regression model using $\widehat{E}(X|W)$ as the covariate,
- 5. adjust standard errors & confidence intervals for estimation of $\widehat{E}(X|W)$.

Fitting a model for repeated error-prone measurements

- Suppose we have data where subjects have W₁ = X + U₁, and some subjects have a second measurement W₂ = X + U₂.
- We assume $X \sim N(\mu_X, Var(X))$ and that $U_1, U_2 \stackrel{\text{iid}}{\sim} N(0, Var(U)).$
- This is the standard one-way analysis of variance or random-intercepts model.
- ► Fitting this model (e.g. using loneway or xtmixed in Stata) gives us estimates of µ_X, Var(X), and Var(U).

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Calculating $\widehat{E}(X|W)$

Under this assumed model:

$$E(X|W) = \mu_X + \frac{Var(X)}{Var(X) + Var(U)}(W - \mu_X)$$

► For those subjects with multiple error-prone measurements, we can use the mean of their error-prone measurements W to predict X:

$${m E}(X|\overline{W})=\mu_X+rac{Var(X)}{Var(X)+Var(U)/k}(\overline{W}-\mu_X)$$

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when \overline{W} is the mean of k measurements.

Stata commands for this simple setting

```
reshape long w, i(id) j(obs)
xtmixed w || id:
predict x_pred, fitted
reshape wide
reg y x_pred
```

Inference and other settings

- The method generalises to the setting of multiple explanatory variables measured with error, and to when we have some explanatory variables measured without error (e.g. gender, treatment group).
- Standard errors and confidence intervals should allow for first stage of estimation, e.g. by using bootstrap methods.
- RC also gives approximately unbiased estimates for some other model types (logistic regression, Cox regression).

References

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