# Unbiased estimation of exposure odds ratios in complete records logistic regression 

Jonathan Bartlett

London School of Hygiene and Tropical Medicine www.missingdata.org.uk

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## Outline

When is the exposure odds ratio unbiased?

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## Setting and question

- Binary outcome $Y$, exposure $X$, confouders $C$.
- In general $X$ and/or $C$ could be continuous, and we can use logistic regression for $Y \mid X, C$.
- We may have missing values in $Y, X$ or one or more components of $C$.
- Let $R=1$ denote complete records and $R=0$ denote incomplete records.
- If we drop those with missing values, when will the complete records analysis (CRA) exposure (log) odds ratio $\hat{\beta}_{X}$ be (asymptotically) unbiased?


## A simplification

- First we assume that there no confounders $C$, and that exposure $X$ is binary.
- We can display the full data as

- Odds ratio $=\frac{a \times d}{b \times c}$


## Outcome dependent missingness

- Suppose missingness is outcome dependent, with

$$
P(R=1 \mid X, Y=y)=k_{y}, y=0,1
$$

- Then the expected complete records table is

- Odds ratio $=\frac{k_{0} a \times k_{1} d}{k_{1} b \times k_{0} c}=\frac{a \times d}{b \times c}$
- i.e. unbiased
- This result of course justifies the validity of odds ratios in case-control studies.


## Exposure dependent missingness

- Now suppose missingness is exposure dependent, with $P(R=1 \mid X=x, Y)=k_{x}, x=0,1$.
- Then the expected complete records table is

- Odds ratio $=\frac{k_{0} a \times k_{1} d}{k_{0} b \times k_{1} c}=\frac{a \times d}{b \times c}$
- i.e. unbiased
- This result (covariate dependent missingness) holds in fact more generally, e.g. for risk ratios and other regression models.


## Confounders

- Now suppose we have a categorical confounder $C$ with levels $c=1, . ., L$.
- Let the full data exposure/outcome table for level c of the confounder be

- The full data odds ratio for the exposure effect at confounder level $C=c$ is $\frac{a_{c} \times d_{c}}{b_{c} \times c_{c}}$
- Assuming no effect modification, we then estimate the odds ratio by pooling across confounder levels, e.g. with Mantel-Haenszel or logistic regression.


## Outcome/confounder dependent missingness

- Now suppose we have missingness dependent on $Y$ and $C$, i.e. $P(R=1 \mid Y=y, X, C=c)=k_{y, c}, y=0,1, c=1, . ., L$.
- The expected complete record table is then

- Odds ratio $=\frac{k_{0, c} a_{c} \times k_{1, c} d_{c}}{k_{1, c} b_{c} \times k_{0, c} c_{c}}=\frac{a_{c} \times d_{c}}{b_{c} \times c_{c}}$.
- i.e. still unbiased


## Exposure/confounder dependent missingness

- Now suppose we have missingness dependent on $X$ and $C$, i.e. $P(R=1 \mid Y, X=x, C=c)=k_{x, c}, x=0,1, c=1, . ., L$.
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- i.e. still unbiased


## Exposure/outcome dependent missingness

- In general if missingness depends on $X$ and $Y$, we will obtain biased estimates of the exposure effect.
- A special case exists however where we still obtain unbiased estimates.
- Specifically if $P(R=1 \mid Y, X, C)=s(X, C) \times t(Y, C)$ for some functions $s(X, C), t(Y, C)$.
- This result can be viewed as applying the previous results in turn.


## Exposure/outcome dependent missingness continued

- This could arise for example if two variables are partially observed with one mechanism depending on $(X, C)$ and the second depending on $(Y, C)$.
- If we let $R_{1}$ and $R_{2}$ denote the observation indicators for the two variables, we need

$$
\begin{aligned}
P(R=1 \mid Y, X, C) & =P\left(R_{1}=1, R_{2}=1 \mid Y, X, C\right) \\
& =P\left(R_{1}=1 \mid X, C\right) \times P\left(R_{2}=1 \mid Y, C\right)
\end{aligned}
$$

## Summary

The results apply more generally, to continuous exposure $X$, multiple confounders $C$, with logistic regression used for estimation.

Letting $\hat{\beta}_{0}, \hat{\beta}_{X}, \hat{\beta}_{C}$ denote the corresponding log odds ratios, we have

| Missingness dependent on | $\hat{\beta}_{0}$ | $\hat{\beta}_{X}$ | $\hat{\beta}_{C}$ |
| :--- | :--- | :--- | :--- |
| Neither $Y, X$ nor $C$ | Unbiased | Unbiased | Unbiased |
| $Y$ | Biased | Unbiased | Unbiased |
| $X, C$ | Unbiased | Unbiased | Unbiased |
| $Y, C$ | Biased | Unbiased | Biased |
| $Y, X, C$ | Biased | Biased* | Biased |

* in general


## MCAR/MAR/MNAR

- So far we have not said which variable(s) have missing values.
- Depending on what type of mechanism is assumed to be in action and which variable(s) have missing values, missingness could be MCAR, MAR or even MNAR.
- e.g. if exposure is partially observed, with missingness dependent on exposure level, and possibly confounders, this is MNAR.
- But so long as missingness is independent of $Y$, given $X$ and $C, \hat{\beta}_{X}$ is unbiased.


## Extensions

- If we include interactions between $X$ and some components of C, the results continue to apply.
- With survival data, if the event rate is low and follow-up similar for subjects, the results also apply approximately due to link between logistic and Cox regression.
- In this case the event indicator takes the place of the binary outcome $Y$.


## Caveat

- The arguments implicitly assume that the outcome model is correctly specified.
- i.e. no effect interactions (if not included).
- Provided any misspecifications are not severe, results should still approximately apply (see illustrative) example.


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- The results depend on how the probability of being a complete record depend on ( $Y, X, C$ ).
- In some cases our study specific knowledge may support one of the sufficient conditions needed for $\beta_{X}$ to be estimated without bias.


## Missingness in a confounder

Divide the confounders into $C=\left(C_{1}, C_{2}\right)$, with $C_{1}$ partially observed

| Missingness in $C_{1}$ <br> found related to | Plausible missingness <br> mechanism | $\hat{\beta}_{X}$ unbiased |
| :---: | :---: | :---: |
| $C_{2}$ | $C_{2}$ | $Y e s$ |
| $X$ and possibly $C_{2}$ | $X$ and $C_{2}$ | $Y$ Yes |
| $Y$ and possibly $C_{2}$ | $Y$ and $C_{2}$ | $Y e s$ |
|  | $X$ and $Y$ | Generally no |
| $X, Y$ and possibly $C_{2}$ | $C$ and $X$ | $Y e s$ |
|  | $C$ and $Y$ | $Y e s$ |

## Missingness in exposure

| Missingness in $X$ <br> found related to | Plausible missingness <br> mechanism | $\hat{\beta}_{X}$ unbiased |
| :---: | :---: | :---: |
| $C$ | $C$ | Yes |
| $Y$ | $Y$ | Yes |
| $C$ and $Y$ | $C$ and $Y$ | $Y e s$ |
|  | $X$ and $Y$ | Generally no |
|  | $X$ and $C$ | $Y e s$ |

## Missingness in outcome

| Missingness in $Y$ <br> found related to | Plausible missingness <br> mechanism | $\hat{\beta}_{X}$ unbiased |
| :---: | :---: | :---: |
| $X$ | $X$ | $Y e s$ |
| $C$ | $C$ | $Y e s$ |
| $X$ and $C$ | $X$ and $C$ | $Y e s$ |
|  | $Y$ and $C$ | $Y e s$ |
|  | $X$ and $Y$ | Generally no |

## Investigating missingness

- So in some situations we may from the data be able to make some tentative conclusions regarding missingness and thus the unbiasedness of $\hat{\beta}_{X}$.
- With missingness in multiple variables things inevitably become more complex.
- Here our substantive knowledge is crucial to judge the plausibility of assumptions.


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## Illustrative example

- We illustrate using data from a cohort study of professional pilots in the UK [1].
- This study's aim was to include all professional flight crew who held a license at some point between 1989 and 1999.
- For our analyses, follow-up starts at recruitment, where various variables of interest were recorded.
- Crew data was linked to UK health registers to obtain vital status up to 2006.


## Outcome model specification

- For this illustrative analysis, we consider a binary outcome $Y$ representing death in the first 15 years of follow-up.
- We consider a exposure 'number of accrued flying hours' at baseline, categorised into $<400$ hours, $400-5499$, and $>5500$ hours.
- Confounders were age at entry, smoking status, BMI and type of route.
- Our analyses here use data from 11,841 male crew, who had complete data at baseline for exposure and confounders.


## Simulation setup

- We first fitted a logistic regression model to the full data.
- For each of 8 missingness mechanisms, we make data missing and perform the complete records analysis.
- For each mechanism we repeat 10,000 times, and present the mean estimates across 10,000 realizations of missingness process.


## Results

Log odds ratios

| Missingness <br> dependent on | $400-5499$ hours <br> vs $<400$ hours | $\geq 5500$ hours <br> vs $<400$ hours |
| :--- | :--- | :--- |
| Full data | 0.55 | 0.59 |
| MCAR | 0.57 | 0.61 |
| Event indicator $(Y)$ | 0.56 | 0.61 |
| Age $(C)$ | 0.51 | 0.54 |
| Age and flying hours $(C, X)$ | 0.52 | 0.56 |
| Event indicator and age $(Y, C)$ | 0.66 | 0.74 |
| Event indicator and flying hours $(Y, X)$ | 1.59 | 2.64 |
| Event indicator and flying hours* $(Y, X)$ | 0.59 | 0.67 |

* $P(R=1 \mid Y, X, C)=s(X) t(Y)$


## Interpretation

- For the most part, results are as expected from theory.
- For missingness dependent on outcome and confounder (age), estimates are somewhat different to full data estimates.
- This is likely due to the fact the outcome model is not exactly correctly specified: the exposure effect varies somewhat by age.
- Missingness dependent on $C$ means complete records have different $C$ distribution, leading to some small bias.


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- Estimates of exposure effect from complete records logistic regression are unbiased under a wide range of missingness assumptions [2].
- Most of the results presented are not new [3, 4, 5], but perhaps are not widely appreciated.
- Directed acyclic graphs can be useful for considering possible missingness mechanisms and thus plausiblity of assumptions [6].


## Summary

- Of course even when CRA is unbiased, it is not efficient.
- If missingness is in confounders, alternative approaches will gain efficiency.
- If the data are missing at random, multiple imputation can be used.
- If a confounder is missing not at random, but independently of outcome, a weighting/imputation approach can be used to improve upon CRA efficiency [7].


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